

## User's Guide for Conducting *A Priori* Power Analyses in SAS

Before we describe how to use our power programs, complete these preliminary steps:

1. Download and save *MIXREGLS.zip* from: <https://www.jstatsoft.org/article/view/v052i12>
2. Extract (aka, unzip) *MIXREGLS.zip* and save all extracted files to a known destination location. You will be required to use this destination location in our power programs (e.g., we extracted the files to: E:/MIXREGLS).
3. Locate the *Calculate\_Values.xlsx* Excel file that will use to calculate the values to enter into the power programs.
4. Locate the *Detect.sas* and *Predict.sas* files.

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### Example Scenario

Say we are interested in quantifying and predicting individual differences in the intra-individual variability of positive affect in a sample of adolescents. The independent variable in this study will be the average hourly moderate-to-vigorous physical activity (MVPA) measured across the study period (a level-2, individual-level predictor). Said another way, we are interested in whether individual differences exist in how much adolescents vary around their own average amount of positive affect and whether these differences can be explained by their average amount of MVPA. Based on evidence observed in the literature, we expect the unconditional intra-class correlation (ICC) for positive affect to be 0.50, assuming that both the outcome-scale random intercept variance ( $\sigma_{b_0}^2$ ) and the residual variance ( $\sigma_e^2$ ) will equal 2. We expect scale-model random intercept variance ( $\sigma_{t_0}^2$ ) to be 0.10 and expect the correlation between the location- and scale-model random intercepts to be 0.30. Finally, we expect the average amount of MVPA ( $M_{\text{centered}} = 0$ ,  $SD = 0.75$ ) to explain 10% of scale-model random intercept variance.

## The Power to Detect Individual Differences in Intra-Individual Variability

Estimating the empirical power to detect individual differences in intra-individual variability is a two-step process: 1) calculate values using the *Calculate\_Values.xlsx* Excel file and then 2) enter these values into and run the *Detect.sas* power program. Open the *Detect* tab of the *Calculate\_Values* Excel file. In this file, we can enter values in any yellow-shaded cell. Further, gray-shaded cells are calculated values for information purposes and are not used in the power program, whereas blue-shaded cells are values that will be used in the power programs.

Based on our example scenario, we enter 2 for both the location-model random intercept variance and residual variance in cells B13 and C13, respectively, enter 0.10 for the scale-model random intercept variance in cell D13, and enter 0.30 for the correlation between the location- and scale-model random intercepts. As shown in Figure 1, the unconditional ICC is calculated automatically for us using two different formulas that yield equivalent results—the ICC in cell E13 does not consider scale-model random intercept variance (cell E13), whereas the ICC in cell F13 does consider scale-model random intercept variance.

A Priori Variance Estimates					
Location-Model Random Intercept Variance $\sigma_{b_0}^2$ (Outcome Scale)	Total Residual Variance $\sigma_e^2$ (Outcome Scale)	Scale-Model Random Intercept Variance $\sigma_{t_0}^2$	Correlation of Location- and Scale-Model Random Intercepts $\rho_{b_0,t_0}$	Unconditional Intra-Class Correlation (ICC)	Unconditional Intra-Class Correlation (ICC)  (Hedeker et al., 2008; Equation 10)
2.000	2.000	0.100	0.300	0.500	0.500

**Figure 1.** Entered values based on the example scenario

Based on the values provided in Figure 1, the Excel file automatically calculates the values to enter into our power program as shown in Figure 2.

Values to Enter into the DetectSMRI Power Program			
Location-Model Random Intercept Variance $\sigma_{b_0}^2$ <b>(b0var)</b>	Scale-Model Fixed Intercept $\tau_0$ <b>(tau0)</b>	Scale-Model Random Intercept Variance $\sigma_{t_0}^2$ <b>(t0var)</b>	Correlation of Location- and Scale-Model Random Intercepts $\rho_{b_0,t_0}$ <b>(b0t0corr)</b>
2.000	0.643	0.100	0.300

**Figure 2.** Values to enter into the *Detect* power program

These values include the location-model random intercept variance on the outcome scale (*b0var* in cell C24;  $\sigma_{b_0}^2$ ), the scale-model fixed intercept on the log scale (*tau0* in cell D24;  $\tau_0$ ), the scale-model random intercept variance (*t0var* in cell E24;  $\sigma_{t_0}^2$ ), and the correlation between the location- and scale-model random intercepts (*b0t0corr* in cell F24;  $\rho_{b_0,t_0}$ ).

After calculating the necessary values in the Excel file, we turn our attention to the *Detect.sas* power program used to detect scale-model random intercept variance. In the *Detect.sas* file, we only need to change the values associated with the objects defined in lines 8-30; we define values by changing quantities after the equal sign. Below, we provide a description of what effect each line specifies as well as the value we use based on our example scenario.

**Line 14** – Enter the file location where we extracted (or unzipped) the MIXREGLS.zip file – here, E:/MIXREGLS.

**Line 16** – Enter the number of individuals ( $N$ ) – here, 50.

**Line 18** – Enter the number of repeated occasions ( $n_i$ ) – here, 15.

**Line 20** – Enter the fixed intercept for the location-model (*beta0*;  $\beta_0$ ) – here, 0. This is an

outcome-scale value that has no effect on estimated power.

**Line 22** – Enter the location-model random intercept variance on the outcome scale ( $b0var; \sigma_{b_0}^2$ )

– here, 2.

**Line 24** – Enter the fixed intercept for the scale-model for the residual variance ( $tau0; \tau_0$ ) –

here, 0.643.

**Line 26** – Enter the scale-model random intercept variance ( $t0var; \sigma_{t_0}^2$ ) – here, 0.10.

**Line 28** – Enter the correlation between the location- and scale-model random intercepts

( $b0t0corr; \rho_{b_0, t_0}$ ) – here, 0.30. It should be noted that MixRegLS does *not* estimate

this correlation directly, but can allow the location-model random intercept to

influence the scale-model random intercept using fixed linear or quadratic effects (see

Hedeker & Nordgren, 2013, for complete details). We calculate this correlation after

the model has been estimated (see lines 289).

**Line 30** – Enter the number of replications we want to use when estimating power. These power

programs can take a long time to run especially as the number of individuals and/or

repeated measures increase, so although we use 1,000 replications for this example,

we recommend starting with 10 replications to get a feel for the power. In general, we

recommend at least 100 replications when determining the final power estimate.

After entering all values, lines 8 through 30 in this SAS program should look like Figure 3

below. Click Run.

```

8 *****
9 ***          Define MACRO Variable Values Below          ***
10 ***          Use Calculate_Values.xlsx to Determine Values      ***
11 *****;
12
13 *Folder Directory for (unzipped) MIXREGLS;
14   %LET MIXREGLS = E:\MIXREGLS;
15 *Individuals;
16   %LET N = 50;
17 *Repeated Occasions;
18   %LET ni = 15;
19 *Location-Model Fixed Intercept (Outcome Scale);
20   %LET beta0 = 0;
21 *Location-Model Random Intercept Variance (Outcome Scale);
22   %LET b0var = 2;
23 *Scale-Model Fixed Intercept (Log Scale);
24   %LET tau0 = 0.643;
25 *Scale-Model Random Intercept Variance;
26   %LET t0var = 0.10;
27 *Correlation between Location- and Scale-Model Random Intercepts;
28   %LET b0t0corr = 0.30;
29 *Number of Replications for Power Analysis;
30   %LET Nreps = 1000;

```

**Figure 3.** Specified macro variables in the *Detect* power program in SAS

In SAS, MixRegLS runs via command prompt (i.e., Windows command processor) that shows the program’s iteration history. After the program completes, results will be outputted to the ODS HTML window presenting both the parameter recovery estimates as shown in Figure 4, as well as the empirical power estimate as shown in Figure 5.

Variable	Label	N	Mean	Lower 95% CL for Mean	Upper 95% CL for Mean	25th Pctl	50th Pctl	75th Pctl
b0var_YS_DS	b0var: LM RI Variance (Outcome Scale) - Yes SMRI	970	1.949	1.922	1.976	1.644	1.927	2.244
evar_YS_DS	evar: Total Residual Variance (Outcome Scale) - Yes SMRI	970	2.002	1.993	2.011	1.903	1.996	2.096
ICC_NS	ICC: Unconditional Intra-Class Correlation - No SMRI	970	0.488	0.485	0.492	0.450	0.490	0.526
ICC_YS	ICC: Unconditional Intra-Class Correlation - Yes SMRI	970	0.488	0.485	0.492	0.450	0.490	0.526
evar_YS	evar: Total Residual Variance (Log Scale) - Yes SMRI	970	0.692	0.687	0.696	0.644	0.691	0.740
b0var_YS	b0var: LM RI Variance (Log Scale) - Yes SMRI	970	0.643	0.629	0.657	0.497	0.656	0.808
tau0_YS	tau0: SM Fixed Intercept (Log Scale) - Yes SMRI	970	0.644	0.639	0.648	0.598	0.644	0.688
t0var_YS	t0var: SM RI Variance	970	0.096	0.093	0.099	0.061	0.091	0.126
b0t0corr_YS	b0t0corr: Correlation b/w Location- and Scale-Model RI	970	0.311	0.295	0.326	0.147	0.321	0.480

**Figure 4.** Parameter Recovery Estimates from the *Detect* power program in SAS

Scale-Model RI (t0var) Significant: 0=No, 1=Yes				
t0var_YS_SS	Frequency	Percent	Cumulative Frequency	Cumulative Percent
0	211	21.75	211	21.75
1	759	78.25	970	100.00

**Figure 5.** Power Estimates from the *Detect* power program in SAS

Although we specified 1,000 replications, the results were based on 970 replications because 30 replications could not be estimated by MixRegLS. We have found this to be a problem when the scale-model random intercept variance is estimated to be  $\sim 0$  and/or when location-model random intercept variance or residual variance are large. In general, these types of errors tend to go by the wayside with larger samples. The results in Figure 7 indicate that based on the model parameters we entered, we would achieve approximately 78% power to detect significant individual differences in intra-individual variability (i.e., scale-model random intercept variance) with 50 adolescents who each have 15 repeated occasions. The 95% Clopper-Pearson confidence interval (not shown) was estimated to be 76% to 81%.

### **The Power to Predict Individual Differences in Intra-Individual Variability**

Once we identify the amount or proportion of scale-model random intercept variance we want to detect, we turn our attention to prediction of this variance component. Similar to the procedure to detect scale-model random intercept variance, estimating the empirical power to predict individual differences in intra-individual variability (i.e., scale-model random intercept variance) is a two-step process. Based on our example scenario, we open the *Predict* tab of the *Calculate\_Values* Excel file and enter 2 for both the location-model random intercept variance and total residual variance in cells B13 and C13, respectively, enter 0.10 for the scale-model random intercept variance in cell D13, and enter 0.30 for the correlation between the location- and scale-model random intercepts as previously shown in Figure 1. We then provide information about the MVPA predictor variable (called  $W_i$ ) as shown in Figure 6.

Information about the Predictor Variable $W_i$					
		Continuous Predictor (Assumed Normally Distributed)		Binary Predictor	
Is Predictor $W_i$ Continuous?  (0=No; 1=Yes)	Proportion of Scale-Model Random Intercept Variance Explained  (Pseudo- $R^2$ )	Mean of Predictor $W_i$  (Grand-Mean-Centered with a Mean of 0)	Variance of Predictor $W_i$	Mean of Predictor $W_i$	Variance of Predictor $W_i$
1	0.100	0	0.563	0.500	0.250

**Figure 6.** Information about scale-model predictor  $X_i$

Specifically, we enter a 1 into cell B26 to indicate that predictor  $W_i$  is continuous and enter 0.10 into cell C26 for pseudo- $R^2$  because we expect this predictor to explain 10% of scale-model random intercept variance. Our power program assumes continuous predictors are grand-mean-centered such that they have a mean of 0. We enter 0.563 into cell D13 to indicate the predictor's variance (i.e.,  $0.75 \times 0.75$ ). For continuous predictors, we can ignore cells F26 and G26 because they apply only to a binary predictor. Had our predictor been binary, however, we would ignore cells D26 and E26 and instead enter the predictor's mean (i.e., the proportion of 1s) into cell F26 upon which its variance would be automatically calculated for us in cell G26. Based on the values provided in Figures 1 and 6, the Excel file automatically calculates the values to enter into our power program as shown in Figure 7.

Values to Enter into the PredictSMRI Power Program								
Location-Model Random Intercept Variance  $\sigma_{b_0}^2$  (b0var)	Scale-Model Fixed Intercept  $\tau_0$  (tau0)	Scale-Model Random Intercept Variance  $\sigma_{\tau_0}^2$  (t0var)	Correlation of Location- and Scale-Model Random Intercepts  $\rho_{b_0\tau_0}$  (b0t0corr)	Unstandardized Scale-Model Fixed Effect for Predictor $W_i$ $\tau_1$  (tau1)	Pseudo- $R^2$  (R2)	Is Predictor Continuous?  (Continuous)	Mean of Predictor $W_i$  (Wmean)	Variance of Predictor $W_i$  (Wvar)
2.000	0.643	0.100	0.300	0.133	0.100	1	0	0.563

**Figure 7.** Values to enter into the *Predict* power program

The values in Figure 7 include the location-model random intercept variance on the outcome scale ( $b0var$  in cell B37;  $\sigma_{b_0}^2$ ), the scale-model fixed intercept ( $tau0$  in cell C37;  $\tau_0$ ), the amount scale-model random intercept variance ( $t0var$  in cell D37;  $\sigma_{t_0}^2$ ), and the correlation between the location- and scale-model random intercepts ( $b0t0corr$  in cell E37;  $\rho_{b_0,t_0}$ ). Also provided is the unstandardized fixed effect for the predictor variable ( $tau1$  in cell F37;  $\tau_1$ ) and population pseudo- $R^2$  ( $R2$  in cell G37). Finally, predictor information including an indicator of whether the predictor is continuous (*Continuous* in cell H37) as well as the predictor's mean and variance ( $Wmean$  in cell I37 and  $Wvar$  in cell J37) are also provided.

After calculating the necessary values in the Excel file, we turn our attention to the *Predict.sas* power program used to predict scale-model random intercept variance. In this SAS file, we only need to change the values associated with the macro variables located in lines 8 to 43. Although most macro variables are identical to those described in the *Detect.sas* program, except for those listed on lines 29-41, below we provide a description of what each line requires as well as the value we use in our example scenario.

**Line 14** – Enter the file location where we extracted (or unzipped) the MIXREGLS.zip file – here, “E:/MIXREGLS/”. Note that the quotation marks as well as the forward slash after the file name are critically important.

**Line 16** – Enter the number of individuals ( $N$ ) – here, 50.

**Line 18** – Enter the number of repeated occasions ( $n_i$ ) – here, 20.

**Line 20** – Enter the fixed intercept for the location-model ( $beta0$ ;  $\beta_0$ ) – here, 0. This is an outcome-scale value that has no effect on estimated power.

**Line 22** – Enter the location-model random intercept variance on the outcome scale ( $b0var$ ;  $\sigma_{b_0}^2$ ) – here, 2.

- Line 24** – Enter the fixed intercept for the scale-model for the residual variance ( $\tau_0$ ) – here, 0.643.
- Line 26** – Enter the scale-model random intercept variance ( $\sigma_{t_0}^2$ ) – here, 0.10.
- Line 28** – Enter the correlation between the location- and scale-model random intercepts ( $\rho_{b_0, t_0}$ ) – here, 0.30. It should be noted that MixRegLS does *not* estimate this correlation directly, but can allow the location-model random intercept to influence the scale-model random intercept using fixed linear or quadratic effects (see Hedeker & Nordgren, 2013, for complete details). We calculate this correlation after the model has been estimated (see lines 457 and 459).
- Line 30** – Enter the unstandardized fixed effect of the scale-model predictor  $X_i$  ( $\tau_1$ ) – here, 0.133.
- Line 32** – Enter the population pseudo- $R^2$  value ( $R^2$ ) – here, 0.10.
- Line 34** – Enter either 0 or 1 to indicate whether the scale-model predictor  $X_i$  is continuous or binary (0 = binary, 1 = continuous) – here, 1.
- Line 36** – Enter the mean of the scale-model predictor  $W_i$  ( $Wmean$ ). Note that continuous predictors are assumed to be grand-mean-centered with a mean of 0 – here, 0.
- Line 41** – Enter the variance of the scale-model predictor  $X_i$  ( $Xvar$ ). Note that, although the variance of a binary predictor is defined by its mean and this value is not used when sampling from the Bernoulli distribution, some non-null value must be entered after the equal sign on lines 41 for the SAS program to execute.
- Line 43** – Enter the number of replications we want to use when estimating power. These power programs can take a long time to run especially as the number of individuals and/or repeated measures increase, so although we use 1,000 replications for this example,

we recommend starting with 10 replications to get a feel for the power. In general, we recommend at least 100 replications when determining the final power estimate.

After entering all values, lines 8 through 31 in this SAS program should look like Figure 8 below. Highlight the entire SAS program and click run.

```
8 *****
9 ***          Define MACRO Variable Values Below          ***
10 ***      Use Calculate_Values.xlsx to Determine Values      ***
11 *****;
12
13 *Folder Directory for (unzipped) MIXREGLS;
14   %LET MIXREGLS = E:\MIXREGLS;
15 *Individuals;
16   %LET N = 50;
17 *Repeated Occasions;
18   %LET ni = 20;
19 *Location-Model Fixed Intercept (Outcome Scale);
20   %LET beta0 = 0;
21 *Location-Model Random Intercept Variance (Outcome Scale);
22   %LET b0var = 2;
23 *Scale-Model Fixed Intercept (Log Scale);
24   %LET tau0 = 0.643;
25 *Scale-Model Random Intercept Variance;
26   %LET t0var = 0.10;
27 *Correlation between Location- and Scale-Model Random Intercepts;
28   %LET b0t0corr = 0.30;
29 *Scale-Model Fixed Predictor W_i Effect;
30   %LET tau1 = 0.133;
31 *Pseudo-R2;
32   %LET R2 = 0.10;
33 *Is the Predictor W_i Continuous? (1=Yes, 0=No);
34   %LET Continuous = 1;
35   *Mean of Predictor W_i (=0 if W_i is continuous);
36   %LET Wmean = 0;
37   *Variance of Continuous Predictor W_i;
38   *Does not apply to binary predictor, but must have
39   non-null, non-zero value after equal sign for program
40   to run;
41   %LET Wvar = 0.563;
42 *Number of Replications for Power Analysis;
43   %LET Nreps = 1000;
```

**Figure 8.** Specified macro variables in the *Predict* power program in SAS

In SAS, MixRegLS runs via command window and results will be outputted to the ODS HTML window presenting both the parameter recovery estimates (as shown in Figure 9) as well as the empirical power estimate shown in Figure 10.

Variable	Label	N	Mean	Lower 95% CL for Mean	Upper 95% CL for Mean	25th Pctl	50th Pctl	75th Pctl
beta0_YP	LM Fixed Intercept (Outcome Scale) - Yes Predictor	994	-0.005	-0.018	0.008	-0.143	-0.004	0.129
b0var_YP_DS	LM RI Variance (Outcome Scale) - Yes Predictor	994	1.968	1.943	1.994	1.686	1.939	2.216
tau0_YP	SM Fixed Intercept (Log Scale) - Yes Predictor	994	0.643	0.639	0.647	0.597	0.643	0.688
tau1_YP	SM Fixed Effect of Predictor W (Log Scale)	994	0.131	0.126	0.136	0.077	0.133	0.182
t0var_NP	SM RI Variance - No Predictor	994	0.097	0.094	0.099	0.067	0.093	0.121
t0var_YP	SM RI Variance - Yes Predictor	994	0.084	0.081	0.086	0.056	0.080	0.108
PseudoR2	Proportion Reduction SM RI Variance from Predictor W	994	0.145	0.135	0.154	0.032	0.099	0.210
b0t0corr_YP	Correlation b/w LM and SM RI - Yes Predictor	994	0.322	0.307	0.336	0.157	0.332	0.480

**Figure 9.** Parameter Recovery Estimates from the *Predict* power program in SAS

Scale-Model Predictor W Significant: 0=No, 1=Yes				
tau1_SS	Frequency	Percent	Cumulative Frequency	Cumulative Percent
0	626	62.98	626	62.98
1	368	37.02	994	100.00

**Figure 10.** Power Estimates from the *Predict* power program in SAS

Although we specified 1,000 replications, the results were based on 994 replications because for 6 replications the scale-model random intercept variance could not be estimated by MixRegLS either before or after adding the predictor variable. The results in Figure 10 indicate that based on the model parameters we entered, we would achieve approximately 37% power to predict 10% of the individual differences in intra-individual variability (i.e., 10% of the scale-model random intercept variance) with 50 adolescents who each have 20 repeated occasions. The 95% Clopper-Pearson confidence interval (not shown) was estimated to be 34% to 40%.

When considering the empirical power estimates from the *Detect* and *Predict* programs, we can see that although we had sufficient power to detect individual differences in intra-individual variability (i.e., scale-model random intercept variance) with 50 adolescents and 20 repeated occasions, power to predict 10% of this variability was woefully inadequate. As such,

we would subsequently re-run the *Predict* program with increased numbers of adolescents and/or repeated occasions to ensure adequate power to predict individual differences in intra-individual variability.